

Asymptotic Properties for the Parameter Estimator in the Near-Explosive Autoregressive Process

Mingming Yu, Jiao Meng

Nanjing University of Aeronautics and Astronautics, Nanjing

Email: mengyilianmeng@163.com, zbmengjiao@sina.com

Received: Oct. 13th, 2014; revised: Nov. 11th, 2014; accepted: Nov. 20th, 2014

Copyright © 2014 by authors and Hans Publishers Inc.

This work is licensed under the Creative Commons Attribution International License (CC BY).

<http://creativecommons.org/licenses/by/4.0/>



Open Access

Abstract

In this paper, we focus our attention on the following near-explosive autoregressive process:

$$\begin{cases} X_{k,n} = \theta_n X_{k-1,n} + \varepsilon_{k,n} \\ \varepsilon_{k,n} = \rho_n \varepsilon_{k-1,n} + V_k \end{cases}$$
. When $\theta_n \rightarrow 1$ and $\rho_n \rightarrow 1$ in the near-explosive case, the asymptotic distributions for the least squares estimator of θ_n can be obtained.

Keywords

Autoregressive Process, Least Squares Estimator, Near-Explosive

近爆炸性自回归序列中参数估计量的渐近性质

于明明, 孟 娇

南京航空航天大学, 南京

Email: mengyilianmeng@163.com, zbmengjiao@sina.com

收稿日期: 2014年10月13日; 修回日期: 2014年11月11日; 录用日期: 2014年11月20日

摘要

本论文的目的是研究近爆炸性自回归序列 $\begin{cases} X_{k,n} = \theta_n X_{k-1,n} + \varepsilon_{k,n} \\ \varepsilon_{k,n} = \rho_n \varepsilon_{k-1,n} + V_k \end{cases}$ 中, 当 $\theta_n \rightarrow 1$, $\rho_n \rightarrow 1$ 时参数 θ_n 最小二乘估计量的渐近分布。

关键词

自回归序列, 最小二乘法估计量, 近爆炸

1. 介绍

本文我们讨论下面自回归序列:

$$\begin{cases} X_{k,n} = \theta_n X_{k-1,n} + \varepsilon_{k,n} \\ \varepsilon_{k,n} = \rho_n \varepsilon_{k-1,n} + V_k \end{cases}, \quad 1 \leq k \leq n, \quad n \geq 1. \quad (1)$$

并且满足下列假设:

- (1) θ_n, ρ_n 是未知参数, 并且满足 $\theta_n = 1 + \frac{\gamma_1}{k_n}$, $\rho_n = 1 + \frac{\gamma_2}{k_n}$, $\gamma_1 > \gamma_2 > 0$, 其中 $k_n \rightarrow +\infty$ 且 $k_n = o(n)$;
- (2) $\{V_k, k = 1, 2, \dots, n\}$ 是独立同分布 (i.i.d.) 随机变量, 且 $EV_k = 0$, $EV_k^2 = \sigma^2$;
- (3) $X_0 = \varepsilon_0 = 0$ 。

为了估计未知参数 θ_n , 我们通过使 $\Delta_n(\theta_n) = \sum_{k=1}^n (X_{k,n} - \theta_n X_{k-1,n})^2$ 的值达到最小得到 θ_n 的最小二乘估计为 $\hat{\theta}_n = \frac{\sum_{k=1}^n X_{k,n} X_{k-1,n}}{\sum_{k=1}^n X_{k-1,n}^2}$, $n \geq 1$ 。

首先, 当 θ_n 是一个固定的常数 θ 且 $\rho_n \equiv 0$ 时, 这种情况下 $\hat{\theta}_n$ 的渐近分布已经被很多科研者证明出来, 我们可以参考文献 [1] [2], 情况再复杂一点, 当 ρ_n 也是一个固定的常数 ρ 时, 在 [3] 中, 作者证出了 $\hat{\theta}_n$ 的渐近正态性。其次, 科研者们不再满足于研究 θ_n 固定时的情况, 假设 θ_n 是一个可变的序列, $\rho_n \equiv 0$, Chan 和 Wei 在 [4] 中证出了在 $\theta_n = 1 + \frac{\gamma_1}{n}$ 时, $\hat{\theta}_n - \theta_n$ 渐近于一个布朗运动, 当 $\rho_n = 1 + \frac{\gamma_2}{n}$, 在 [5] 中, 作者给出了 $\hat{\theta}_n - \theta_n$ 的渐近分布。

最后, 在这篇文章中我们来考虑 $\theta_n = 1 + \frac{\gamma_1}{k_n}$, $\rho_n = 1 + \frac{\gamma_2}{k_n}$, $\gamma_1 > \gamma_2 > 0$ 时 $\hat{\theta}_n - \theta_n$ 的渐近性质。我们有如下主要的结论:

定理 1: 当 n 充分大时

$$\frac{\gamma_1 + \gamma_2}{2\gamma_1(\gamma_1 - \gamma_2)} k_n \theta_n \rho_n^{-n} (\hat{\theta}_n - \theta_n) \xrightarrow{L} \frac{\xi_\rho}{\xi_\theta},$$

这里 \xrightarrow{L} 表示依分布收敛, $(\xi_\theta, \xi_\rho) \sim N(0, \Lambda)$, 其中 $\Lambda = \begin{pmatrix} \frac{\sigma^2}{2\gamma_1} & \frac{\sigma^2}{\gamma_1 + \gamma_2} \\ \frac{\sigma^2}{\gamma_1 + \gamma_2} & \frac{\sigma^2}{2\gamma_2} \end{pmatrix}$ 。

2. 定理的证明

下面为了计算方便, 对所有的 $1 \leq l \leq n$, 我们令

$$\begin{aligned} P_{l,n} &= \sum_{k=1}^l X_{k,n} X_{k-1,n}, \quad S_{l,n} = \sum_{k=1}^l X_{k,n}^2, \\ M_{l,n} &= \sum_{k=1}^l X_{k-1,n} V_k, \quad N_{l,n} = \sum_{k=2}^l X_{k-2,n} V_k, \end{aligned}$$

计 $P_{n,n} = P_n$, $S_{n,n} = S_n$, $M_{n,n} = M_n$, $N_{n,n} = N_n$, 这样我们得到

$$\widehat{\theta}_n - \theta_n = \frac{P_n - \theta_n S_{n-1,n}}{S_{n-1,n}}.$$

另外, 定义一些新的序列

$$\begin{aligned} \xi_{\theta_n} &= \frac{1}{\sqrt{k_n}} \sum_{l=1}^n \theta_n^{-l} V_l, \quad \eta_{\theta_n} = \frac{1}{\sqrt{k_n}} \sum_{l=1}^n \theta_n^{-(n-l)-1} V_l, \\ \xi_{\rho_n} &= \frac{1}{\sqrt{k_n}} \sum_{l=1}^n \rho_n^{-l} V_l, \quad \eta_{\rho_n} = \frac{1}{\sqrt{k_n}} \sum_{l=1}^n \rho_n^{-(n-l)-1} V_l. \end{aligned}$$

为了证明定理内容, 我们引入下面引理

引理 2.1: 对于模型(1), 我们得到

$$X_{n,n}^2 = \frac{\theta_n^2}{(\theta_n - \rho_n)^2} \theta_n^{2n} k_n \xi_{\theta_n}^2 + \frac{\rho_n^2}{(\theta_n - \rho_n)^2} \rho_n^{2n} k_n \xi_{\rho_n}^2 - \frac{2\theta_n \rho_n}{(\theta_n - \rho_n)^2} \theta_n^n \rho_n^n k_n \xi_{\theta_n} \xi_{\rho_n}$$

$$\text{以及 } X_{n,n} \varepsilon_{n,n} = \frac{\theta_n}{(\theta_n - \rho_n)} \theta_n^n \rho_n^n k_n \xi_{\theta_n} \xi_{\rho_n} - \frac{\rho_n}{\theta_n - \rho_n} \rho_n^{2n} k_n \xi_{\rho_n}^2, \quad \varepsilon_{n,n}^2 = k_n \rho_n^{2n} \xi_{\rho_n}^2.$$

证明: 由 Phillips 和 Magdalinos [6], 我们知 $X_{k,n} = \sum_{l=1}^k \theta_n^{k-l} \varepsilon_{l,n}$, $\varepsilon_{k,n} = \sum_{l=1}^k \rho_n^{k-l} V_{l,n}$, 故对所有的 $1 \leq k \leq n$, 有

$$X_{k,n} = \frac{\theta_n}{\theta_n - \rho_n} \theta_n^k \sum_{l=1}^k \theta_n^{-l} V_l - \frac{\rho_n}{\theta_n - \rho_n} \rho_n^k \sum_{l=1}^k \rho_n^{-l} V_l, \quad (2)$$

通过简单的计算, 我们可以得到

$$\begin{aligned} X_{n,n}^2 &= \frac{\theta_n^2}{(\theta_n - \rho_n)^2} \theta_n^{2n} \left(\sum_{l=1}^n \theta_n^{-l} V_l \right)^2 + \frac{\rho_n^2}{(\theta_n - \rho_n)^2} \rho_n^{2n} \left(\sum_{l=1}^n \rho_n^{-l} V_l \right)^2 - \frac{2\theta_n \rho_n}{(\theta_n - \rho_n)^2} \theta_n^n \rho_n^n \sum_{l=1}^n \theta_n^{-l} V_l \sum_{l=1}^n \rho_n^{-l} V_l \\ &= \frac{\theta_n^2}{(\theta_n - \rho_n)^2} \theta_n^{2n} k_n \xi_{\theta_n}^2 + \frac{\rho_n^2}{(\theta_n - \rho_n)^2} \rho_n^{2n} k_n \xi_{\rho_n}^2 - \frac{2\theta_n \rho_n}{(\theta_n - \rho_n)^2} \theta_n^n \rho_n^n k_n \xi_{\theta_n} \xi_{\rho_n}, \end{aligned}$$

另外, 通过公式(2)和 $\varepsilon_{n,n} = \rho_n^n \sum_{l=1}^n \rho_n^{-l} V_l$, 我们可以得到

$$X_{n,n} \varepsilon_{n,n} = \frac{\theta_n}{(\theta_n - \rho_n)} \theta_n^n \rho_n^n k_n \xi_{\theta_n} \xi_{\rho_n} - \frac{\rho_n}{\theta_n - \rho_n} \rho_n^{2n} k_n \xi_{\rho_n}^2 \text{ 以及 } \varepsilon_{n,n}^2 = k_n \rho_n^{2n} \xi_{\rho_n}^2.$$

引理 2.2: 当 n 无穷大时, 我们得到

$$(\xi_{\theta_n}, \eta_{\theta_n}, \xi_{\rho_n}, \eta_{\rho_n})^T \xrightarrow{L} (\xi_\theta, \eta_\theta, \xi_\theta, \eta_\theta)^T, \quad \text{其中 } (\xi_\theta, \eta_\theta, \xi_\theta, \eta_\theta)^T \sim N(0, \Sigma),$$

并且

$$\Sigma = \begin{pmatrix} \frac{\sigma^2}{2\gamma_1} & 0 & \frac{\sigma^2}{\gamma_1 + \gamma_2} & 0 \\ 0 & \frac{\sigma^2}{2\gamma_1} & 0 & \frac{\sigma^2}{\gamma_1 + \gamma_2} \\ \frac{\sigma^2}{\gamma_1 + \gamma_2} & 0 & \frac{\sigma^2}{2\gamma_2} & 0 \\ 0 & \frac{\sigma^2}{\gamma_1 + \gamma_2} & 0 & \frac{\sigma^2}{2\gamma_2} \end{pmatrix}$$

证明：通过[7]中的推论5.5，我们只需要证明：对任意的非零向量 $v = (v_1, v_2, v_3, v_4)$ ，有 $v(\xi_{\theta_n}, \eta_{\theta_n}, \xi_{\rho_n}, \eta_{\rho_n})^T \xrightarrow{L} v(\xi_\theta, \eta_\theta, \xi_\theta, \eta_\theta)^T$ ，对任意的 $1 \leq l \leq n$ ，令

$$\xi_{nl} = \frac{1}{\sqrt{k_n}} (v_1 \theta_n^{-l} + v_2 \theta_n^{-(n-l)-1} + v_3 \rho_n^{-l} + v_4 \rho_n^{-(n-l)-1}) V_l, \text{ 有 } v(\xi_{\theta_n}, \eta_{\theta_n}, \xi_{\rho_n}, \eta_{\rho_n})^T = \sum_{l=1}^n \xi_{nl}.$$

一方面，因为 $\{\xi_{nl}, 1 \leq l \leq n\}$ 是一个独立非同分布随机变量列，通过一些简单的计算可得到

$$E\left(\sum_{l=1}^n \xi_{nl}\right)^2 = v E\left(\left(\xi_{\theta_n}, \eta_{\theta_n}, \xi_{\rho_n}, \eta_{\rho_n}\right)^T \left(\xi_{\theta_n}, \eta_{\theta_n}, \xi_{\rho_n}, \eta_{\rho_n}\right)\right) v^T = v \Sigma_n v^T \rightarrow v \Sigma v^T$$

其中

$$\Sigma_n = \frac{\sigma^2}{k_n} \sum_{l=1}^n \begin{pmatrix} \theta_n^{-2l} & \theta_n^{-n-1} & (\theta_n \rho_n)^l & \rho_n^{-n-1} (\theta_n / \rho_n)^{-l} \\ \theta_n^{-n-1} & \theta_n^{-2n-2+2l} & \theta_n^{-n-1} (\theta_n / \rho_n)^l & (\theta_n \rho_n)^{-n-1+l} \\ (\theta_n \rho_n)^l & \theta_n^{-n-1} (\theta_n / \rho_n)^l & \rho_n^{-2l} & \rho_n^{-n-1} \\ \rho_n^{-n-1} (\theta_n / \rho_n)^{-l} & (\theta_n \rho_n)^{-n-1+l} & \rho_n^{-n-1} & \rho_n^{-2n-2+2l} \end{pmatrix}$$

另一方面，因为当 n 无穷大时，存在一个足够大的数 K ，有

$$\frac{8}{k_n} \sum_{l=1}^n (v_1^2 \theta_n^{-2l} + v_2^2 \theta_n^{-2n-2+2l} + v_3^2 \rho_n^{-2l} + v_4^2 \rho_n^{-2n-2+2l}) \rightarrow \frac{4(v_1^2 + v_2^2)}{\gamma_1} + \frac{4(v_3^2 + v_4^2)}{\gamma_2} \leq K$$

并且

对任意的 $\delta > 0$ ，我们有

$$\begin{aligned} & \sum_{l=1}^n E(\xi_{nl}^2) I_{\{|\xi_{nl}| > \delta\}} \\ & \leq \frac{8}{k_n} \sum_{l=1}^n \left((v_1^2 \theta_n^{-2l} + v_2^2 \theta_n^{-2n-2+2l} + v_3^2 \rho_n^{-2l} + v_4^2 \rho_n^{-2n-2+2l}) \cdot E\left(V_l^2 I_{\{8(v_1^2 \theta_n^{-2l} + v_2^2 \theta_n^{-2(n-l)-2} + v_3^2 \rho_n^{-2l} + v_4^2 \rho_n^{-2(n-l)-2}) V_l^2 > \delta^2 k_n\}}\right)\right) \end{aligned}$$

因为 $\theta_n > 1$, $\rho_n > 1$ 故有 $8(v_1^2 \theta_n^{-2l} + v_2^2 \theta_n^{-2(n-l)-2} + v_3^2 \rho_n^{-2l} + v_4^2 \rho_n^{-2(n-l)-2}) < 8(v_1^2 + v_2^2 + v_3^2 + v_4^2)$ ，所以我们可得

$$E\left(V_l^2 I_{\{8(v_1^2 \theta_n^{-2l} + v_2^2 \theta_n^{-2(n-l)-2} + v_3^2 \rho_n^{-2l} + v_4^2 \rho_n^{-2(n-l)-2}) V_l^2 > \delta^2 k_n\}}\right) \leq E\left(V_1^2 I_{\{8(v_1^2 + v_2^2 + v_3^2 + v_4^2) V_1^2 > \delta^2 k_n\}}\right)$$

即有 $\sum_{l=1}^n E(\xi_{nl}^2) I_{\{|\xi_{nl}| > \delta\}} \leq K E\left(V_1^2 I_{\left\{V_1^2 > \frac{\delta^2 k_n}{8(v_1^2 + v_2^2 + v_3^2 + v_4^2)}\right\}}\right)$ ，又因为 V_1^2 的可积性及 $k_n \rightarrow \infty$ ，所以有

$$E \left(V_1^2 I_{\left\{ V_1^2 > \frac{\delta^2 k_n}{8(v_1^2 + v_2^2 + v_3^2 + v_4^2)} \right\}} \right) \rightarrow 0, \text{ 因此我们可以得到 Lindeberg 条件 } \sum_{l=1}^n E \left(\xi_{nl}^2 I_{\{|\xi_{nl}| > \delta\}} \right) \rightarrow 0;$$

基于上面两方面的原因，我们可以得到这个引理的证明。

引理 2.3: 当 n 无穷大时，我们可以得到

$$\frac{\sum_{k=1}^n \varepsilon_{k-1,n} V_k}{k_n \rho_n^n} = \xi_{\rho_n} \eta_{\rho_n} + o_p(1), \quad \frac{\varepsilon_{n,n}^2}{k_n \rho_n^{2n}} = \xi_{\rho_n}^2, \quad \frac{X_{n,n}^2}{k_n^3 \theta_n^{2n}} = \frac{1}{(\gamma_1 - \gamma_2)^2} \xi_{\theta_n}^2 + o_p(1), \text{ 以及}$$

$$\frac{M_n}{k_n^2 \theta_n^n} = \frac{1}{\gamma_1 - \gamma_2} \xi_{\theta_n} \eta_{\theta_n} + o_p(1), \quad \frac{X_{n,n} \varepsilon_{n,n}}{k_n^2 \theta_n^n \rho_n^n} = \frac{1}{\gamma_1 - \gamma_2} \xi_{\theta_n} \xi_{\rho_n} + o_p(1).$$

证明: 首先，我们通过 Phillip 和 Magdalions 在[6]中对公式(10)的证明以及引理 2.1 可以得到

$$\frac{\sum_{k=1}^n \varepsilon_{k-1,n} V_k}{k_n \rho_n^n} = \xi_{\rho_n} \eta_{\rho_n} + o_p(1) \text{ 以及 } \frac{\varepsilon_{n,n}^2}{k_n \rho_n^{2n}} = \xi_{\rho_n}^2.$$

其次，通过引理 2.1，引理 2.2 以及 $\frac{\rho_n^n}{\theta_n^n} = o(1)$ ，我们可以得到

$$\frac{X_{n,n}^2}{k_n^3 \theta_n^{2n}} = \frac{\theta_n^2}{(\theta_n - \rho_n)^2 k_n^2} \xi_{\theta_n}^2 + \frac{\rho_n^{2n+2}}{(\theta_n - \rho_n)^2 k_n^2 \theta_n^{2n}} \xi_{\rho_n}^2 - \frac{2\theta_n \rho_n^{n+1}}{(\theta_n - \rho_n)^2 k_n^2 \theta_n^n} \xi_{\theta_n} \xi_{\rho_n} = \frac{1}{(\gamma_1 - \gamma_2)^2} \xi_{\theta_n}^2 + o_p(1)$$

另外

$$\begin{aligned} \frac{M_n}{k_n^2 \theta_n^n} &= \frac{\theta_n}{(\theta_n - \rho_n) k_n^2} \sum_{k=2}^n \theta_n^{-(n-k)-1} V_k \cdot \sum_{l=1}^n \theta_n^{-l} V_l - \frac{\rho_n^n}{\theta_n^n} \cdot \frac{\rho_n}{(\theta_n - \rho_n) k_n^2} \sum_{k=2}^n \rho_n^{-(n-k)-1} V_k \cdot \sum_{l=1}^n \rho_n^{-l} V_l \\ &\quad - \frac{\theta_n}{(\theta_n - \rho_n) k_n^2 \theta_n^n} \sum_{k=2}^n \theta_n^{k-1} V_k \cdot \sum_{l=k}^n \theta_n^{-l} V_l + \frac{\rho_n^n}{\theta_n^n} \cdot \frac{\theta_n}{(\theta_n - \rho_n) k_n^2 \rho_n^n} \sum_{k=2}^n \rho_n^{k-1} V_k \cdot \sum_{l=k}^n \rho_n^{-l} V_l, \end{aligned}$$

易知：

$$\frac{\theta_n}{(\theta_n - \rho_n) k_n^2} \sum_{k=2}^n \theta_n^{-(n-k)-1} V_k \cdot \sum_{l=1}^n \theta_n^{-l} V_l = \frac{1}{\gamma_1 - \gamma_2} \xi_{\theta_n} \eta_{\theta_n} + o_p(1) \quad (3)$$

$$\frac{\rho_n}{(\theta_n - \rho_n) k_n^2} \sum_{k=2}^n \rho_n^{-(n-k)-1} V_k \cdot \sum_{l=1}^n \rho_n^{-l} V_l = \frac{1}{\gamma_1 - \gamma_2} \xi_{\rho_n} \eta_{\rho_n} + o_p(1) \quad (4)$$

又因为

$$\frac{\theta_n^{-n}}{k_n^2 (\theta_n - \rho_n)} \sum_{k=2}^n V_k^2 = O_p \left(\theta_n^{-n} \frac{n}{k_n} \right) = o_p(1), \quad E \left(\frac{\theta_n^{1-n}}{k_n^2 (\theta_n - \rho_n)} \sum_{k=2}^n \theta_n^{k-1} V_k \cdot \sum_{l=k+1}^n \theta_n^{-l} V_l \right)^2 = o(1)$$

故

$$\frac{\theta_n}{(\theta_n - \rho_n) k_n^2 \theta_n^n} \sum_{k=2}^n \theta_n^{k-1} V_k \cdot \sum_{l=k}^n \theta_n^{-l} V_l = \frac{\theta_n^{-n}}{k_n^2 (\theta_n - \rho_n)} \sum_{k=2}^n V_k^2 + \frac{\theta_n^{1-n}}{k_n^2 (\theta_n - \rho_n)} \sum_{k=2}^n \theta_n^{k-1} V_k \cdot \sum_{l=k+1}^n \theta_n^{-l} V_l = o_p(1) \quad (5)$$

同理可知：

$$\frac{\theta_n}{(\theta_n - \rho_n) k_n^2 \rho_n^n} \sum_{k=2}^n \rho_n^{k-1} V_k \cdot \sum_{l=k}^n \rho_n^{-l} V_l = o_p(1) \quad (6)$$

结合(3)(4)(5)(6)得:

$$\frac{M_n}{k_n^2 \theta_n^n} = \frac{1}{\gamma_1 - \gamma_2} \xi_{\theta_n} \eta_{\theta_n} + o_p(1).$$

最后, 易知 $\frac{X_{n,n} \varepsilon_{n,n}}{k_n^2 \theta_n^n \rho_n^n} = \frac{\theta_n}{k_n(\theta_n - \rho_n)} \xi_{\theta_n} \xi_{\rho_n} - \frac{\rho_n^{n+1}}{k_n \theta_n^n (\theta_n - \rho_n)} \xi_{\rho_n}^2 = \frac{1}{\gamma_1 - \gamma_2} \xi_{\theta_n} \xi_{\rho_n} + o_p(1)$.

引理 2.4: 当 n 无穷大时, 我们有 $\frac{S_{n-1,n}}{k_n^4 \theta_n^{2n}} = \frac{1}{2\gamma_1(\gamma_1 - \gamma_2)^2} \xi_{\theta_n}^2 + o_p(1)$.

证明: 由[3]中(A.14)和(A.23)式, 令 $L_n = \sum_{l=1}^n V_l^2$, 得到

$$\begin{aligned} & \left(1 - (\theta_n + \rho_n)^2 - (\theta_n \rho_n)^2\right) S_{n-1,n} \\ &= -X_{n,n}^2 - (\theta_n \rho_n)^2 X_{n-1,n}^2 + L_n - 2\theta_n \rho_n (\theta_n + \rho_n) P_{n-1,n} + 2(\theta_n + \rho_n) M_n - 2\theta_n \rho_n N_n; \end{aligned} \quad (7)$$

又

$$P_n = \frac{\theta_n + \rho_n}{1 + \theta_n \rho_n} S_{n-1,n} + \frac{1}{1 + \theta_n \rho_n} M_n + \frac{\theta_n \rho_n}{1 + \theta_n \rho_n} X_{n,n} X_{n-1,n}, \quad (8)$$

并且

$$N_n = \frac{M_n - \sum_{k=1}^n \varepsilon_{k-1,n} V_k}{\theta_n}, \quad X_{n-1,n}^2 = \frac{X_{n,n}^2 + \varepsilon_{n,n}^2 - 2X_{n,n} \varepsilon_{n,n}}{\theta_n}, \quad X_{n,n} X_{n-1,n} = \frac{X_{n,n}^2 - X_{n,n} \varepsilon_{n,n}}{\theta_n},$$

因此根据(7)式我们可以得到 $S_{n-1,n} = \frac{1}{\theta_n^2 - 1} X_{n,n}^2 + R_{n1}$, 其中

$$\begin{aligned} R_{n1} &= \frac{2\theta_n \rho_n}{(1 - \theta_n \rho_n)(\theta_n^2 - 1)} X_{n,n} \varepsilon_{n,n} - \frac{\rho_n^2 (1 + \theta_n \rho_n)}{(1 - \theta_n \rho_n)(1 - \theta_n^2)(1 - \rho_n^2)} \varepsilon_{n,n}^2 + \frac{2\theta_n}{(1 - \theta_n \rho_n)(1 - \theta_n^2)} M_n \\ &+ \frac{1 + \theta_n \rho_n}{(1 - \theta_n \rho_n)(1 - \theta_n^2)(1 - \rho_n^2)} L_n + \frac{2\rho_n (1 + \theta_n \rho_n)}{(1 - \theta_n \rho_n)(1 - \theta_n^2)(1 - \rho_n^2)} \sum_{k=1}^n \varepsilon_{k-1,n} V_k, \end{aligned}$$

由引理 2.3 得 $\frac{1}{\theta_n^2 - 1} X_{n,n}^2 = \frac{k_n^4 \theta_n^{2n}}{2\gamma_1(\gamma_1 - \gamma_2)^2} \xi_{\theta_n}^2 + o_p(k_n^4 \theta_n^{2n})$ 及 $R_{n1} = o_p(k_n^4 \theta_n^{2n})$, 引理得证。

定理 1 的证明

由(7)(8)式得, $P_n - \theta_n S_{n-1,n} = -\frac{\rho_n}{1 - \theta_n \rho_n} X_{n,n} \varepsilon_{n,n} - \frac{\rho_n^3}{(1 - \theta_n \rho_n)(1 - \rho_n^2)} \varepsilon_{n,n}^2 + R_{n2}$,

其中 $R_{n2} = \frac{\rho_n}{(1 - \theta_n \rho_n)(1 - \rho_n^2)} L_n + \frac{1}{1 - \theta_n \rho_n} M_n + \frac{2\rho_n^2}{(1 - \theta_n \rho_n)(1 - \rho_n^2)} \sum_{k=1}^n \varepsilon_{k-1,n} V_k$.

由引理 2.1 和引理 2.3 得

$$-\frac{\rho_n}{1 - \theta_n \rho_n} X_{n,n} \varepsilon_{n,n} - \frac{\rho_n^3}{(1 - \theta_n \rho_n)(1 - \rho_n^2)} \varepsilon_{n,n}^2 = \frac{k_n^3 \theta_n^n \rho_n^n}{\gamma_1^2 - \gamma_2^2} \xi_{\theta_n} \xi_{\rho_n} + o_p(k_n^3 \theta_n^n \rho_n^n) \quad (9)$$

又经过简单的计算得

$$R_{n2} = o_p(k_n^3 \theta_n^n \rho_n^n), \quad (10)$$

结合(9)(10), 定理得证。

致 谢

本论文是在我与同学孟娇的合作中完成的。感谢导师给予我们的支持，感谢南京航空航天大学数学系的各位老师给予我们的指导和帮助，感谢各位文献作者的成果给予我们的借鉴。

参考文献 (References)

- [1] Anderson, T.W. (1959) On asymptotic distributions of estimators of parameters of stochastic difference equations. *Annals of Mathematical Statistics*, **32**, 676-687.
- [2] Zhou, Z.Z. and Lin, Z.Y. (1958) Asymptotic theory for LAD estimation of moderate deviations from a unit root. *Statistics and Probability Letters*, **29**, 1188-1197.
- [3] Bercu, B. and Proia, F. (2013) A sharp analysis on the asymptotic behavior of the Durbin-Watson statistic for the first-order autoregressive process. *ESAIM: Probability and Statistics*, **17**, 500-530.
- [4] Chan, N.H. and Wei, C.Z. (1987) Asymptotic inference for nearly nonstationary AR(1) processes. *Annals of Statistics*, **15**, 1050-1063.
- [5] Nabeya, S. and Perron, P. (1994) Local asymptotic distributions related to the AR(1) model with dependent errors. *Journal of Econometrics*, **62**, 229-264.
- [6] Phillips, P.C.B. and Magdalinos, T. (2007) Limit theory for moderate deviations from a unit root. *Journal of Econometrics*, **136**, 115-130.
- [7] Kallenberg, O. (2002) Foundations of modern probability. Springer, Berlin.